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A Gated Recurrent Unit Aided Lee-Carter Model for mortality projection

A.U Shelleng¹, H.G Dikko², J. Garba³, M. Abdulkarim⁴ and B.B Alhaji⁵

Department of Mathematics, Gombe State University, Gombe¹

Department of Statistics, Ahmadu Bello University Zaria, Kaduna^{2,3}

Department of Computer Science, Ahmadu Bello University Zaria, Kaduna⁴

Department of Mathematics, Nigerian Defence Academy, Kaduna⁵

ABSTRACT

Over the past decades, new models for mortality modelling have been developed. A pioneering and one of the most influential model, among others, is the Lee-Carter (LC) model, developed by Lee and Carter in 1992. The method traditionally used in the LC model shows obvious drawback in describing the shape of future mortality. This work used a more robust approach based on a deep learning technique that differs from the literature. More precisely, a Gated recurrent unit (GRU) was integrated with the Lee-Carter model to increase its capacity for prediction. The LC-GRU model performs well in terms of predicted accuracy. The strength of the gated recurrent unit network lies in its capability to handle the non-linearity inherent in mortality data and long-term dependencies in time series data. The research results shows that the LC-GRU model has minimum errors and thus performs better than the LC model.

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1. Introduction

The need to forecast into the future and have a good understanding of what future mortality rates and life expectancy will look like is of major importance to researchers. Numerous methods have been developed recently to project mortality using stochastic models (McNown and Rogers, 1989; Lee and Carter, 1992; Alho, 1990; Currie *et al.*, 2004; Cairns *et al.*, 2006; Renshaw and Haberman, 2006). The Lee-Carter model, created by Lee and Carter in 1992, is one of the most often used models for projecting mortality rates among the several models that have been used to forecast mortality. The model has garnered a lot of interest and has come to be regarded as a benchmark for mortality modeling and projection. For instance, the Lee-Carter model was applied to Japan (Wilmoth, 1993),

Australia (Booth *et al.*, 2002), Austria (Carter and Prskawetz, 2001), and Belgium (Tuljapurkar, 2008). (Brouhns *et al.*, 2002).

The Lee-Carter model (LC model) has a number of shortcomings despite its widespread acceptance (Lee, 2000). These shortcomings have been criticized by researchers such as Gutterman and Vanderhoof (1999). Researchers and actuaries have started using more advanced methods for predicting future mortality rates in order to treat longevity risk and predict future mortality rates with more accuracy. With recent improvements in computational ability, artificial intelligence (AI) are back on stage. AI techniques enclose numerous algorithms that learn from vast datasets with numerous attributes, and they are highly effective in spotting obscure and difficult-to-detect patterns.

The lack of wide used of AI approaches in demography is largely attributable to what are frequently referred to as "black boxes," which means that it is impossible to fully comprehend the inner workings of the hidden layer and that these workings are thus challenging to interpret. Additionally, the algorithms utilized in AI are based on data rather than specific hypothesis, which are of greater interest to demographers.

To correct some of the drawbacks of the existing models which includes inability to handle non-linearity inherent in mortality data and long-term dependence in time series, this work combines the Lee-Carter model with a Gated recurrent unit (GRU). The GRU makes it possible to get mortality projections that are more consistent with the variability and nonlinear trends observed in mortality. More specifically, the GRU is set up to handle lengthy data sequences, creating a memory that can hold onto important links between data. In this sense, the time series framework of the GRU enables the prediction of future mortality over time, taking into account the significant influence of the historical mortality pattern and accurately reproducing it in the trend that is predicted. The information will also be preserved over

time via the GRU, preventing it from gradually disappearing throughout processing.

2. Material and Methods

The data used in this research work were obtained from the United Nation World Population Prospect (WPP) database. Simulated data from Gompertz distribution and Age-specific mortality rates records of Nigeria male adults within the ages of 15 to 59 from 2009 to 2020 were used.

2.1 Gated Recurrent Unit aided Lee-Carter model (LC-GRU)

The main components of the Lee-Carter approach are the base model and the Time series component. In the first stage, a_x is estimated using Ordinary least square method, b_x and k_t are estimated Singular value decomposition method. In the second stage, the Gated recurrent unit model is used to model and extrapolate the fitted values of k_t . Thus, if we let the mortality data collected at times to be $g(1), g(2), \dots, g(J)$ then the LC-GRU model is given as:

$$\ln(m_{xg(t)}) = a_x + b_x k_{g(t)} + \varepsilon_{xg(t)} \quad (1)$$

$$m_{xg(t)} = \exp^{a_x + b_x k_{g(t)} + \varepsilon_{xg(t)}} \quad (2)$$

$$\text{where } \varepsilon_{xg(t)} \sim N(0, \sigma^2), \quad k_{g(t)} = f(k_{g(t-1)}, k_{g(t-2)}, \dots, k_{g(t-J)}) + \varepsilon_{g(t)} \quad (3)$$

$k_{g(t)}$ is a GRU model that approximates a function f that links $k_{g(t)}$ to its time lags, and $j \in N$ represents the number of time lags considered and ε_t is the error term.

$m_{xg(t)}$ represent segmented mortality that is infant mortality, adult mortality or old age mortality, respectively.

a_x is the average shape of the age parameter for segmented mortality.

b_x is the speed of change parameter for segmented mortality.

2.2 Parameter estimation

The least square solution of Equation (1) gives $\hat{a}_x$ as the mean or average of the observed $\ln(m_{xt})$ while $\hat{b}_x$ and $\hat{k}_t$ are the first term of the SVD of the matrix $\ln(m_{xt}) - \hat{a}_x$.

$$\begin{aligned}\ln(m_{xg(t)}) &= a_x + b_x k_{g(t)} + \varepsilon_{g(t)} \\ \varepsilon_{xg(t)} &= \ln(m_{xg(t)}) - a_x - b_x k_{g(t)} \\ \varepsilon_{xg(t)}^2 &= (\ln(m_{xg(t)}) - a_x - b_x k_{g(t)})^2 \\ \sum_{t=1}^T \varepsilon_{xg(t)}^2 &= \sum_{t=1}^T (\ln(m_{xg(t)}) - a_x - b_x k_{g(t)})^2 \\ &= \sum_{t=1}^T (\ln(m_{xt}) - a_x - b_x k_{g(t)})^2 \\ -\sum_{t=1}^T \ln(m_{xg(t)}) + T a_x + b_x \sum_{t=1}^T k_{g(t)} &= 0\end{aligned}$$

To achieve a unique solution we will impose the following restriction $\sum_{t=1}^T k_{g(t)} = 0$

$$\begin{aligned}\therefore T \hat{a}_x &= \sum_{t=1}^T \ln(m_{xg(t)}) \\ \hat{a}_x &= \frac{\sum_{t=1}^T \ln(m_{xg(t)})}{T}\end{aligned}\tag{4}$$

2.3 Singular Value Decomposition

Bell and Monsell introduced the Singular Value Decomposition (SVD) (1991). It is a method for breaking down a matrix into its several constituent

matrices, revealing many of the valuable and intriguing characteristics of the original matrix. Any real $m \times n$ matrix Z can be uniquely decomposed as:

$$Z = B \Sigma K^T\tag{5}$$

The matrices B and K are both defined to be orthogonal matrices. Matrix Σ is a diagonal matrix. The elements along the diagonal of Σ are known as the singular values of the matrix Z . The columns of

B are known as the left-singular vectors. The right-singular vectors are the columns of K .

To obtain parameters b_x and k_t , singular value decomposition (SVD) is applied on the matrix Z where, Z is given as

$$Z = \ln(\hat{m}_{x,t}) - \hat{a}_x \quad (6)$$

That is;

$$Z_{xt} = \begin{pmatrix} \ln(\hat{m}_{1,1}) - \hat{a}_1 & \ln(\hat{m}_{1,2}) - \hat{a}_1 & \ln(\hat{m}_{1,3}) - \hat{a}_1 & \dots & \ln(\hat{m}_{1,n}) - \hat{a}_1 \\ \ln(\hat{m}_{2,1}) - \hat{a}_2 & \ln(\hat{m}_{2,2}) - \hat{a}_2 & \ln(\hat{m}_{2,3}) - \hat{a}_2 & \dots & \ln(\hat{m}_{2,n}) - \hat{a}_2 \\ \ln(\hat{m}_{3,1}) - \hat{a}_3 & \ln(\hat{m}_{3,2}) - \hat{a}_3 & \ln(\hat{m}_{3,3}) - \hat{a}_3 & \dots & \ln(\hat{m}_{3,n}) - \hat{a}_3 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \ln(\hat{m}_{n,1}) - \hat{a}_n & \ln(\hat{m}_{n,2}) - \hat{a}_n & \ln(\hat{m}_{n,3}) - \hat{a}_n & \dots & \ln(\hat{m}_{n,n}) - \hat{a}_n \end{pmatrix}$$

Here we decomposed matrix Z_{xt} into a new set of factors that are optimal and orthogonal based on some criterion.

Thus, the real matrix Z of order $m \times n$ can be uniquely decomposed as:

$$SVD(Z_{x,t}) = B \Sigma K^T \quad (7)$$

where B is $m \times m$ orthogonal matrix that represents the age component.

Σ is $m \times n$ diagonal matrix that represents the singular values.

K is $n \times n$ orthogonal matrix that represents the time component.

$$Z_{xt} = B\Sigma K^T = \begin{pmatrix} b_{11} & b_{12} & b_{13} & \dots & b_{1m} \\ b_{21} & b_{22} & b_{23} & \dots & b_{2m} \\ b_{31} & b_{32} & b_{33} & \dots & b_{3m} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ b_{m1} & b_{m2} & b_{m3} & \dots & b_{mm} \end{pmatrix} \begin{pmatrix} \sigma_{11} & 0 & 0 & \dots & 0 \\ 0 & \sigma_{22} & 0 & \dots & 0 \\ 0 & 0 & \sigma_{33} & \dots & 0 \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ 0 & 0 & 0 & 0 & \sigma_{mn} \end{pmatrix} \begin{pmatrix} k_{11}^T & k_{12}^T & k_{13}^T & \dots & k_{1n}^T \\ k_{21}^T & k_{22}^T & k_{23}^T & \dots & k_{2n}^T \\ k_{31}^T & k_{32}^T & k_{33}^T & \dots & k_{3n}^T \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ k_{n1}^T & k_{n2}^T & k_{n3}^T & \dots & k_{nn}^T \end{pmatrix}$$

Here, $b_{11}, \dots, b_{mm}$ is the Eigen vectors of $Z_{xt} Z_{xt}^T$. $\sigma_1, \dots, \sigma_r$ are the singular values of $Z_{xt} Z_{xt}^T$

satisfying $\sigma_1 \geq \sigma_2 \geq \sigma_3 \geq \dots, \sigma_r > 0$ and $k_1, k_2, \dots, k_n$ are the Eigen vectors of $Z_{xt}^T Z_{xt}$. Here,

Singular values are defined as the positive square roots of the Eigen values of $Z_{xt} Z_{xt}^T$

Matrix $Z_{xt}^T Z_{xt}$ of size $m \times n$ is real and of rank r . Also r of its Eigen values σ_i^2 , $i = 1(1)r$ are positive and therefore real, while the remaining $(n - r)$ are zero.

Theorem: Let $Z_{xt} = B\Sigma K^T$ for $i = 1(1)m$; $j = 1(1)n$ be singular value decomposition of an $m \times n$ real matrix of rank. Then

$$Z_{xt} = B\Sigma K^T = \sum_{i=1}^r b_i \sigma_i k_i^T$$

Proof

$$Z_{xt} = B\Sigma K^T$$

$$Z_{xt} = \begin{pmatrix} \sigma_1 b_{11} & \sigma_2 b_{12} & \sigma_3 b_{13} & \dots & \sigma_r b_{1r} \\ \sigma_1 b_{21} & \sigma_2 b_{22} & \sigma_3 b_{23} & \dots & \sigma_r b_{2r} \\ \sigma_1 b_{31} & \sigma_2 b_{32} & \sigma_3 b_{33} & \dots & \sigma_r b_{3r} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ \sigma_1 b_{r1} & \sigma_2 b_{r2} & \sigma_3 b_{r3} & \dots & \sigma_r b_{rr} \end{pmatrix} \begin{pmatrix} k_{11}^T & k_{12}^T & k_{13}^T & \dots & k_{1r}^T \\ k_{21}^T & k_{22}^T & k_{23}^T & \dots & k_{2r}^T \\ k_{31}^T & k_{32}^T & k_{33}^T & \dots & k_{3r}^T \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ k_{r1}^T & k_{r2}^T & k_{r3}^T & \dots & k_{rr}^T \end{pmatrix}$$

$$\begin{aligned}
Z_{xt} &= \begin{pmatrix} (\sigma_1 b_{11} k_{11}^T + \dots + \sigma_r b_{1r} k_{r1}^T) & (\sigma_1 b_{11} k_{12}^T + \dots + \sigma_r b_{1r} k_{r2}^T) & \dots & (\sigma_1 b_{11} k_{1r}^T + \dots + \sigma_r b_{1r} k_{rr}^T) \\ (\sigma_1 b_{21} k_{11}^T + \dots + \sigma_r b_{2r} k_{r1}^T) & (\sigma_1 b_{21} k_{12}^T + \dots + \sigma_r b_{2r} k_{r2}^T) & \dots & (\sigma_1 b_{21} k_{1r}^T + \dots + \sigma_r b_{2r} k_{rr}^T) \\ \vdots & \vdots & \ddots & \vdots \\ (\sigma_1 b_{r1} k_{11}^T + \dots + \sigma_r b_{rr} k_{r1}^T) & (\sigma_1 b_{r1} k_{12}^T + \dots + \sigma_r b_{rr} k_{r2}^T) & \dots & (\sigma_1 b_{r1} k_{1r}^T + \dots + \sigma_r b_{rr} k_{rr}^T) \end{pmatrix} \\
&= \sigma_1 b_1 k_1^T + \begin{pmatrix} (\sigma_2 b_{12} k_{21}^T + \dots + \sigma_r b_{1r} k_{r1}^T) & (\sigma_2 b_{12} k_{22}^T + \dots + \sigma_r b_{1r} k_{r2}^T) & \dots & (\sigma_2 b_{12} k_{2r}^T + \dots + \sigma_r b_{1r} k_{rr}^T) \\ (\sigma_2 b_{22} k_{21}^T + \dots + \sigma_r b_{2r} k_{r1}^T) & (\sigma_2 b_{22} k_{22}^T + \dots + \sigma_r b_{2r} k_{r2}^T) & \dots & (\sigma_2 b_{22} k_{2r}^T + \dots + \sigma_r b_{2r} k_{rr}^T) \\ \vdots & \vdots & \ddots & \vdots \\ (\sigma_2 b_{r2} k_{21}^T + \dots + \sigma_r b_{rr} k_{r1}^T) & (\sigma_2 b_{r2} k_{22}^T + \dots + \sigma_r b_{rr} k_{r2}^T) & \dots & (\sigma_2 b_{r2} k_{2r}^T + \dots + \sigma_r b_{rr} k_{rr}^T) \end{pmatrix} \\
&= \sigma_1 b_1 k_1^T + \sigma_2 b_2 k_2^T + \dots + \sigma_r b_r k_r^T \\
&= \sigma_1 b_1 k_1^T + \sigma_2 b_2 k_2^T + \dots + \sigma_r b_r k_r^T = \sum_{i=1}^r \sigma_i b_i k_i^T \\
Z_{xt} &= \sum_{i=1}^r \sigma_i b_i k_i^T
\end{aligned}$$

2.4 Merits of the Gate recurrent unit aided LC Model

- i. The number of parameters employed is modest.
- ii. It can handle both the linear and non-linear nature of mortality trends, thus adding confidence to extrapolations.
- iii. The model can handle the long-term dependence associated with time series data.
- iv. The model performs better than other models in terms of its prediction errors

and generates reasonable estimates of mortality forecast.

3. Results and Discussion

The mortality index k_t is modelled using a gated recurrent unit in order to take care of the non-linearity inherent in mortality data and long-term dependence in time series data. As usual, the model parameters b_x and k_t were estimated using SVD.

$$\hat{a}_x = \begin{pmatrix} -1.51258 \\ -1.42149 \\ -1.31180 \\ -1.24766 \\ -1.04828 \\ -1.20401 \\ -1.39770 \\ -1.61356 \\ -1.54164 \end{pmatrix} \quad \hat{b}_x = \begin{pmatrix} -0.22072 \\ -0.58489 \\ -0.07869 \\ 0.34606 \\ 0.27998 \\ -0.06534 \\ -0.15025 \\ -0.37868 \\ -0.46810 \end{pmatrix} \quad \hat{k}_t = \begin{pmatrix} -2.21144 \\ 1.91964 \\ -0.35911 \\ 0.64749 \\ -0.23533 \\ 0.20401 \\ 2.01923 \\ 1.02975 \\ 0.07101 \\ -1.68809 \\ -2.31428 \\ 0.91710 \end{pmatrix}$$

The column matrices $\hat{a}_x$, $\hat{b}_x$ and $\hat{k}_t$ above are the estimated values of the model parameters using SVD. The time component $\hat{k}_t$ is then modelled using both Autoregressive Integrated Moving Average (ARIMA) as used by the Original LC model and a Gated Recurrent Unit (GRU).

Table 1. Forecasted index $\hat{k}_t$ using ARIMA model and GRU for Simulated data

Year	2021	2022	2023	2024	2025	2026	2027
$\hat{k}_t(Arima)$	1.768146	1.908826	2.333063	2.527571	2.794256	2.989057	3.198530
$\hat{k}_t(GRU)$	-0.006259	-0.014376	-0.023553	-0.037314	-0.039273	-0.041222	-0.042646

The Result in table 1 shows the forecasted values of the mortality index $\hat{k}_t$ using both ARIMA and gated recurrent unit.

Table 2. Comparison between $\hat{k}_t$ (ARIMA) and $\hat{k}_t$ (GRU) for simulated data

Method	RMSE	MAE	MAPE
$\hat{k}_t(ARIMA)$	0.3125211	0.2149233	1.279095
$\hat{k}_t(GRU)$	0.2818981	0.2083371	1.168281

It can be seen from table 2, that the GRU models the mortality index k_t better than the traditional ARIMA model used by Lee and Carter in 1992. Projected $\hat{k}_t$ using GRU will then be used to predict the future central death rate m_x .

Table 3 Comparison between LC-ARIMA and LC-GRU for simulated data

Method	RMSE	MAE	MAPE
LC-ARIMA	0.1960865	0.1613628	1.185264
LC-GRU	0.1924219	0.1588562	1.190274

It is observed from table 3 that using the simulated data the (LC-GRU) model performs better than the traditional LC model used by Lee and Carter in 1992. In general, for all the error measures LC-GRU model tend to have lower values compared to the Original LC model.

Real data was also used to model mortality rates using both original LC model the GRU assisted LC model.

The model parameters a_x , b_x and k_t were estimated using the SVD.

$$\hat{a}_x = \begin{pmatrix} -5.0984 \\ -4.9295 \\ -4.8682 \\ -4.8247 \\ -4.71038 \\ -4.5094 \\ -4.3111 \\ -4.0274 \\ -3.7094 \end{pmatrix} \quad \hat{b}_x = \begin{pmatrix} 0.1112 \\ 0.13889 \\ 0.1219 \\ 0.1096 \\ 0.1102 \\ 0.1084 \\ 0.1055 \\ 0.1006 \\ 0.0933 \end{pmatrix} \quad \hat{k}_t = \begin{pmatrix} 0.71673 \\ 1.47743 \\ 0.71673 \\ 0.11885 \\ 0.13909 \\ 0.16763 \\ 0.32585 \\ 0.63814 \\ -0.19014 \\ -0.67994 \\ -1.20958 \\ -2.22082 \end{pmatrix}$$

The above column matrices are the estimated values of the parameters a_x , b_x and k_t using SVD. The time component k_t is then modelled using both Autoregressive Integrated Moving Average (ARIMA) as used by the original LC model and a Gated Recurrent Unit (GRU).

Table 4. Forecasted index $\hat{k}_t$ using ARIMA model and GRU for real data

Year	2021	2022	2023	2024	2025	2026	2027
$\hat{k}_t(Arima)$	-2.552026	-2.660501	-2.696029	-2.707666	-2.711477	-2.712725	-2.71313
$\hat{k}_t(GRU)$	-0.9610381	-1.0214071	-0.951606	-0.7492634	-0.708148	-0.6479220	-0.594411

The Result in table 4 shows the forecasted values of the mortality index k_t using both ARIMA and gated recurrent unit.

Table 5. Comparison between $\hat{k}_t$ (ARIMA) and $\hat{k}_t$ (GRU) for real data

Sex	Method	RMSE	MAE	MAPE
Male	$\hat{k}_t(ARIMA)$	2.809482	2.66587	8.866765
	$\hat{k}_t(GRU)$	1.237017	1.092629	2.74663

It can be seen in table 5 that the GRU model has the minimum errors and thus, models the mortality index k_t better than the traditional ARIMA model used by Lee and Carter in 1992. Projected $\hat{k}_t$ using GRU will then be used predict the future central death rate m_x .

Table 6. Comparison between LC-ARIMA and LC-GRU for real data

Sex	Method	RMSE	MAE	MAPE
Male	LC-ARIMA	0.00427314	0.003826728	0.2578544
	LC-GRU	0.001359604	0.0009497948	0.09406077

It can be seen from table 5 that the (LC-GRU) model has the minimum error measures thus performs better than the traditional LC model used by Lee and Carter in 1992. In general, for all the error measures LC-GRU model tend to have lower values compared to the LC model.

4. Conclusion

After obtaining all the parameters of the model, the performance of the original LC model and the GRU aided approach (GRU-LC model), was assessed. Some of the hyper parameters of the GRU used to improve the performance of the models include: epoch size (100), activation function (Swish), learning rate (0.0001), and the number of neurons in the hidden state. To assess and compare the performance of the original LC model and the LC-GRU model, error measures (RMSE, MAE, and MAPE) were used as measures of accuracy for the age-specific forecasts for the mortality rates. The forecast accuracy of the different approaches is investigated. The summaries in Tables 3, and 6 indicate that different methods show different forecast accuracy in terms of error measures for male, subpopulations and the simulated data. LC-GRU has lower errors as compared to the original LC-model proposed by Lee and Carter in 1992. Therefore LC-GRU performs better than the original LC model.

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